Lecture On-Line Bayesian Tracking

EE 570: Location and Navigation

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Objective

Sequentially estimate on-line the states of a system as it changes over time using observations that are corrupted with noise.

- Filtering: the time of the estimate coincides with the last measurement.
- *Smoothing*: the time of the estimate is within the span of the measurements.
- Prediction: the time of the estimate occurs after the last available measurement.

1 Problem

Given State-Space Equations

$$\vec{\boldsymbol{x}}_{k} = \mathbf{f}_{k}(\vec{\boldsymbol{x}}_{k-1}, \vec{\boldsymbol{w}}_{k-1}) \tag{1}$$

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$$\vec{z}_k = \mathbf{h}_k(\vec{x}_k, \vec{v}_k) \tag{2}$$

where \vec{x}_k is $(n \times 1)$ state vector at time k, \mathbf{f}_k and \mathbf{h}_k are possibly non-linear function $\mathbf{f}_k : \mathfrak{R}^n \times \mathfrak{R}^{n_w} \mapsto \mathfrak{R}^n$ and $\mathbf{h}_k : \mathfrak{R}^m \times \mathfrak{R}^{n_v} \mapsto \mathfrak{R}^m$, respectively, and \vec{w}_k and \vec{v}_k i.i.d state noise. The state process is Markov chain, i.e., $p(\vec{x}_k | \vec{x}_1, \dots, \vec{x}_{k-1}) = p(\vec{x}_k | \vec{x}_{k-1})$ and the distribution of \vec{z}_k conditional on the state \vec{x}_k is independent of previous state and measurement values, i.e., $p(\vec{z}_k | \vec{x}_{1:k-1}) = p(\vec{z}_k | \vec{x}_k)$

Objective

Probabilistically estimate \vec{x}_k using previous measurement $\vec{z}_{1:k}$. In other words, construct the pdf $p(\vec{x}_k | \vec{z}_{1:k})$.

Optimal MMSE Estimate

$$\mathbb{E}\{\|\vec{x}_{k} - \hat{\vec{x}}_{k}\|^{2} |\vec{z}_{1:k}\} = \int \|\vec{x}_{k} - \hat{\vec{x}}_{k}\|^{2} p(\vec{x}_{k} | \vec{z}_{1:k}) d\vec{x}_{k}$$
(3)

in other words find the conditional mean

$$\hat{\vec{x}}_k = \mathbb{E}\{\vec{x}_k | \vec{z}_{1:k}\} = \int \vec{x}_k p(\vec{x}_k | \vec{z}_{1:k}) d\vec{x}_k$$
(4)

2 Bayesian Estimation

3 Kalman Filter

Assumptions

• \vec{w}_k and \vec{v}_k are drawn from a Gaussian distribution, uncorrelated have zero mean and statistically independent.

$$\mathbb{E}\{\vec{\boldsymbol{w}}_{k}\vec{\boldsymbol{w}}_{i}^{T}\} = \begin{cases} \mathbf{Q}_{k} & i=k\\ 0 & i\neq k \end{cases}$$
(5)

$$\mathbb{E}\{\vec{\boldsymbol{v}}_k \vec{\boldsymbol{v}}_i^T\} = \begin{cases} \mathbf{R}_k & i = k\\ 0 & i \neq k \end{cases}$$
(6)

$$\mathbb{E}\{\vec{\boldsymbol{w}_{i}}\vec{\boldsymbol{v}_{i}}^{T}\} = \begin{cases} 0 \quad \forall i,k \end{cases}$$
(7)

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Assumptions

• \mathbf{f}_k and \mathbf{h}_k are both linear, e.g., the state-space system equations may be written as

$$\vec{x}_{k} = \Phi_{k-1} \, \vec{x}_{k-1} + \vec{w}_{k-1}$$
 (8)

$$\vec{y}_k = \mathbf{H}_k \, \vec{x}_k + \vec{v}_k$$
 (9)

where Φ_{k-1} is $(n \times n)$ transition matrix relating \vec{x}_{k-1} to \vec{x}_k , \mathbf{H}_k is $(m \times n)$ matrix provides noiseless connection between measurement and state vectors.

State-Space Equations

$$\hat{\vec{x}}_{k|k-1} = \Phi_{k-1}\hat{\vec{x}}_{k-1|k-1}$$
 (10)

$$\mathbf{P}_{k|k-1} = \mathbf{Q}_{k-1} + \mathbf{\Phi}_{k-1} \mathbf{P}_{k-1|k-1} \mathbf{\Phi}_{k-1}^T$$
(11)

$$\hat{\vec{x}}_{k|k} = \hat{\vec{x}}_{k|k-1} + \mathbf{K}_k \ (\vec{z}_k - \mathbf{H}_k \hat{\vec{x}}_{k|k-1})$$
 (12)

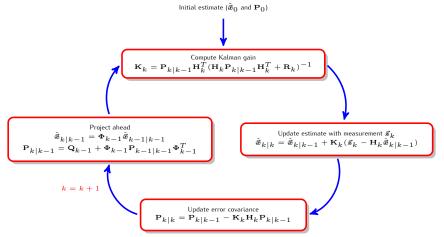
$$\mathbf{P}_{k|k} = (\mathbf{I} - \mathbf{K}_k \mathbf{H}_k) \mathbf{P}_{k|k-1} \tag{13}$$

where \mathbf{K}_k is $(n \times m)$ Kalman gain, and $(\vec{z}_k - \mathbf{H}_k \hat{\vec{x}}_{k|k-1})$ is the measurement innovation.

Kalman Gain

$$\mathbf{K}_{k} = \mathbf{P}_{k|k-1} \mathbf{H}_{k}^{T} (\mathbf{H}_{k} \mathbf{P}_{k|k-1} \mathbf{H}_{k}^{T} + \mathbf{R}_{k})^{-1}$$
(14)

Kalman filter data flow



Sequential Processing

If R is a block matrix, i.e., $R = diag(R^1, R^2, ..., R^r)$. The R^i has dimensions $p^i \times p^i$. Then, we can sequentially process the measurements as:

For i = 1, 2, ..., r

$$\mathbf{K}^{i} = \mathbf{P}^{i-1} (\mathbf{H}^{i})^{T} (\mathbf{H}^{i} \mathbf{P}^{i-1} (\mathbf{H}^{i})^{T} + \mathbf{R}^{i})^{-1}$$
(15)

$$\hat{\vec{x}}_{k|k}^{i} = \hat{\vec{x}}_{k|k}^{i} + \mathbf{K}^{i}(\vec{z}_{k}^{i} - \mathbf{H}^{i}\hat{\vec{x}}_{k|k}^{i-1})$$
(16)

$$\mathbf{P}^{i} = (\mathbf{I} - \mathbf{K}^{i} \mathbf{H}^{i}) \mathbf{P}^{i-1}$$
(17)

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where $\hat{\vec{x}}_{k|k}^{0} = \hat{\vec{x}}_{k|k-1}$, $\mathbf{P}^{0} = \mathbf{P}_{k|k-1}^{0}$ and \mathbf{H}^{i} is $p^{i} \times n$ corresponding to the rows of \mathbf{H} corresponding the measurement being processed.

Observability

The system is observable if the observability matrix

$$\mathcal{O}(k) = \begin{bmatrix} \mathbf{H}(k-n+1) \\ \mathbf{H}(k-n-2)\mathbf{\Phi}(k-n+1) \\ \vdots \\ \mathbf{H}(k)\mathbf{\Phi}(k-1)\dots\mathbf{\Phi}(k-n+1) \end{bmatrix}$$
(18)

where n is the number of states, has a rank of n. The rank of O is a binary indicator and does **not** provide a measure of how close the system is to being unobservable, hence, is prone to numerical ill-conditioning.

A Better Observability Measure

In addition to the computation of the rank of $\mathcal{O}(k)$, compute the Singular Value Decomposition (SVD) of $\mathcal{O}(k)$ as

$$\mathcal{O} = U\Sigma V^* \tag{19}$$

and observe the diagonal values of the matrix Σ . Using this approach it is possible to monitor the variations in the system observability due to changes in system dynamics.

Remarks

- Kalman filter is optimal under the aforementioned assumptions,
- and it is also an unbiased and minimum variance estimate.
- If the Gaussian assumptions is not true, Kalman filter is biased and not minimum variance.
- Observability is dynamics dependent.
- The error covariance update may be implemented using the *Joseph form* which provides a more stable solution due to the guaranteed symmetry.

$$\boldsymbol{P}_{k|k} = \left(\boldsymbol{I} - \boldsymbol{K}_{k} \boldsymbol{H}_{k}\right) \boldsymbol{P}_{k|k-1} \left(\boldsymbol{I} - \boldsymbol{K}_{k} \boldsymbol{H}_{k}\right)^{T} + \boldsymbol{K}_{k} \boldsymbol{R}_{k} \boldsymbol{K}_{k}^{T}$$
(20)

System Model

$$\dot{\vec{x}}(t) = \mathbf{F}(t)\vec{x}(t) + \mathbf{G}(t)\vec{w}(t)$$
(21)

To obtain the state vector estimate $\hat{\vec{x}}(t)$

$$\mathbb{E}\{\dot{\vec{x}}(t)\} = \frac{\partial}{\partial t}\hat{\vec{x}}(t) = \mathbf{F}(t)\hat{\vec{x}}(t)$$
(22)

Solving the above equation over the interval $t - \tau_s, t$

$$\hat{\vec{x}}(t) = e^{\left(\int_{t-\tau_s}^t \mathbf{F}(t')dt'\right)} \hat{\vec{x}}(t-\tau_s)$$
(23)

where \mathbf{F}_{k-1} is the average of \mathbf{F} at times t and $t - \tau_s$.

System Model Discretization

As shown in the Kalman filter equations the state vector estimate is given by

$$\hat{ec{x}}_{k|k-1} = \mathbf{\Phi}_{k-1}\hat{ec{x}}_{k-1|k-1}$$

Therefore,

$$\mathbf{\Phi}_{k-1} = e^{\mathbf{F}_{k-1}\tau_s} \approx \mathbf{I} + \mathbf{F}_{k-1}\tau_s \tag{24}$$

where \mathbf{F}_{k-1} is the average of \mathbf{F} at times t and $t - \tau_s$, and first order approximation is used.

Discrete Covariance Matrix \mathbf{Q}_k

Assuming white noise, small time step, ${\bf G}$ is constant over the integration period, and the trapezoidal integration

$$\mathbf{Q}_{k-1} \approx \frac{1}{2} \left[\mathbf{\Phi}_{k-1} \mathbf{G}_{k-1} \mathbf{Q}(t_{k-1}) \mathbf{G}_{k-1}^T \mathbf{\Phi}_{k-1}^T + \mathbf{G}_{k-1} \mathbf{Q}(t_{k-1}) \mathbf{G}_{k-1}^T \right] \tau_s$$
(25)

where

$$\mathbb{E}\{\vec{w}(\eta)\vec{w}^{T}(\zeta)\} = \mathbf{Q}(\eta)\delta(\eta - \zeta)$$
(26)

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4 EKF

Linearized System

$$\mathbf{F}_{k} = \left. \frac{\partial \mathbf{f}(\vec{x})}{\partial \vec{x}} \right|_{\vec{x} = \hat{\vec{x}}_{k|k-1}}, \qquad \mathbf{H}_{k} = \left. \frac{\partial \mathbf{h}(\vec{x})}{\partial \vec{x}} \right|_{\vec{x} = \hat{\vec{x}}_{k|k-1}}$$
(27)

where

$$\frac{\partial \mathbf{f}(\vec{x})}{\partial \vec{x}} = \begin{pmatrix} \frac{\partial f_1}{\partial x_1} & \frac{\partial f_1}{\partial x_2} & \cdots \\ \frac{\partial f_2}{\partial x_1} & \frac{\partial f_2}{\partial x_2} & \cdots \\ \vdots & \ddots & \vdots \end{pmatrix}, \qquad \frac{\partial \mathbf{h}(\vec{x})}{\partial \vec{x}} = \begin{pmatrix} \frac{\partial h_1}{\partial x_1} & \frac{\partial h_1}{\partial x_2} & \cdots \\ \frac{\partial h_2}{\partial x_1} & \frac{\partial h_2}{\partial x_2} & \cdots \\ \vdots & \ddots & \vdots \end{pmatrix}$$
(28)

5 Example

First Order Markov Noise

State Equation

$$\dot{b}(t) = -\frac{1}{T_c}b(t) + w(t)$$
(29)

Autocorrelation Function

$$\mathbb{E}\{b(t)b(t+\tau)\} = \sigma_{BI}^2 e^{-|\tau|/T_c}$$
(30)

where

$$\mathbb{E}\{w(t)w(t+\tau)\} = Q(t)\delta(t-\tau)$$
(31)

$$Q(t) = \frac{2\sigma_{BI}^2}{T_c} \tag{32}$$

and T_c is the correlation time.

Discrete First Order Markov Noise

State Equation

$$b_k = e^{-\frac{1}{T_c}\tau_s} b_{k-1} + w_{k-1} \tag{33}$$

System Covariance Matrix

$$Q = \sigma_{BI}^2 [1 - e^{-\frac{2}{T_c}\tau_s}]$$
(34)

Autocorrelation of 1st order Markov

Small Correlation Time $T_c = 0.01\,$

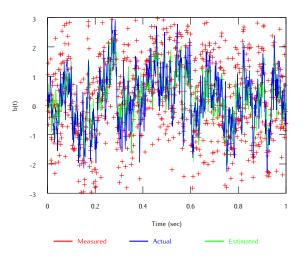
 $R_b(\tau)$

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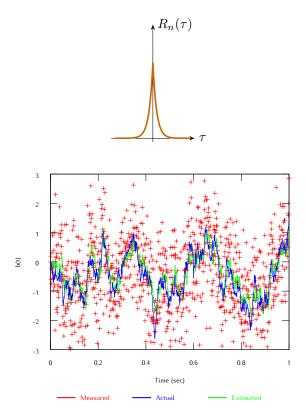


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Larger Correlation Time $T_c = 0.1$



6 Other Solutions

Unscented Kalman Filter (UKF)

Propagates carefully chosen sample points (using unscented transformation) through the true non-linear system, and therefore captures the posterior mean and covariance accurately to the second order.

Particle Filter

A Monte Carlo based method. It allows for a complete representation of the state distribution function. Unlike EKF and UKF, particle filters do not require the Gaussian assumptions.

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7 References

Bayesian Filtering: From Kalman Filters to Particle Filters, and Beyond, by Zhe Chen